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Personality predictors of cognitive dispersion:

A coordinated analysis of data from seven international studies of older adults

Abstract

Objectives: Dispersion in cognitive test performance within a single testing session is proposed as an early marker of poor brain health. Existing research, however, has not investigated factors that may explain individual differences in cognitive dispersion. We investigate the extent to which the Big Five personality traits are associated with cognitive dispersion in older adulthood.

Method: To promote transparency and reliability, we applied pre-registration and conceptual replication via coordinated analysis. Drawing data from seven longitudinal studies of aging ($N_{\text{total}}=33,581$; M_{age} range=56.4-71.2), cognitive dispersion scores were derived from cognitive test results. Independent linear regression models were fit in each study to examine personality traits as predictors of dispersion scores, adjusting for mean cognitive performance and socio-demographics (age, sex, education). Results from individual studies were synthesized using random effects meta-analyses.

Results: Synthesized results revealed that openness was positively associated with cognitive dispersion (0.028, 95% CI:[0.003,0.054]). There was minimal evidence for associations between cognitive dispersion and the other personality traits in independent analyses or meta-analyses. Mean cognitive scores were negatively associated with cognitive dispersion across the majority of studies, while socio-demographic variables were not consistently associated with cognitive dispersion.

Discussion: Higher levels of openness were associated with greater cognitive dispersion across seven independent samples, indicating that individuals higher in openness had more dispersion across cognitive tests. Further research is needed to investigate mechanisms that may help to explain the link between openness and cognitive dispersion, as well as to identify additional individual factors, beyond personality traits, that may be associated with cognitive dispersion.

Keywords: Intra-visit cognitive variability, personality, older adults, coordinated analysis

Total Words: 5769

Key Points

Question: Are personality traits associated with individual differences in cognitive dispersion (i.e., relative variation in performance across cognitive tasks)?

Findings: Across seven independent samples of older adults, individuals higher in openness to experience had greater cognitive dispersion.

Importance: The knowledge that inconsistent cognitive profiles may be characteristic of individuals high in openness, and that associations between other personality traits and cognitive dispersion are likely to be small or null, may better equip medical practitioners to evaluate healthy versus unhealthy cognitive functioning in older adulthood.

Next Steps: Future research should investigate mechanisms that may help to explain the link between openness and cognitive dispersion, as well as additional individual factors that may contribute to cognitive dispersion.

Author Note. Study analytic code (<https://osf.io/5gxa3/>) and pre-registration (<https://osf.io/h8jn5/>), which includes research materials and links to open or public data, are available on the Open Science Framework.

Introduction

Cognitive aging research has traditionally focused on the study of individual differences in cognitive function; specifically, this literature emphasizes investigation of mean population differences or within-person changes in performance over time in one or more cognitive domains. The concept of cognitive dispersion, instead, refers to the degree of relative within-person variation in performance across cognitive tasks assessing various cognitive domains at the same testing occasion (Hilborn et al., 2009; Hultsch et al., 2002). Specifically, computation of cognitive dispersion captures an individual's relative strengths and weaknesses across cognitive tests within a neuropsychological test battery, which may provide a more sensitive index of cognitive ability compared to composite scores based on central tendency. Indeed, existing research suggests that a more uniform cognitive profile represents better cognitive health (Christensen et al., 1999; Hilborn et al., 2009), which is consistent with neuroimaging research indicating that higher white matter integrity is associated with less intraindividual variability across a neuropsychological battery (Halliday et al., 2019). A deeper understanding of the extent to which individual factors, such as personality traits, predict cognitive dispersion may contribute to further understanding of the dynamics between personality traits and cognition, and how personality is involved in the cognitive aging process.

An important body of literature suggests that cognitive dispersion may be an early marker of poor brain health, dementia and mortality. That is, while some cognitive dispersion is normal, a high degree may represent inefficient cognitive processing, beyond performance on any individual neuropsychological test (Holtzer et al., 2008; Michael Malek-Ahmadi et al., 2018). In support of this idea, various publications have reported an association between cognitive dispersion and neuropathology in cortical (Bielak et al., 2010; Bunce et al., 2013; Das et al., 2014; Fjell et al., 2011) and neocortical (Bangen et al., 2019) areas, as well as with

increased levels of amyloid beta (Duchek et al., 2009) and neurofibrillary tangles in healthy individuals and in individuals with mild cognitive impairment and AD dementia (Malek-Ahmadi et al., 2017). Furthermore, higher baseline cognitive dispersion scores predict progression to MCI and dementia (Holtzer et al., 2008; Roalf et al., 2016; Tales et al., 2012) and may be similarly or independently sensitive to early pathological change compared with APOEε4 carrier status, as well as measures of hippocampal atrophy and cerebral spinal fluid (Anderson et al., 2016; Gleason et al., 2018).

This accumulated evidence has led to the postulation that cognitive dispersion may be a valuable index to identify individuals at increased risk of poor brain health for selection into trials and interventions aiming to delay the onset or reduce the risk of cognitive deterioration. Importantly, estimation of within-person cognitive dispersion based on neuropsychological procedures is simple and cost-effective for clinicians (Holtzer et al., 2008). In particular, the adoption of these indices would repurpose familiar and currently available neuropsychological tools to potentially optimize their sensitivity and specificity for dementia detection (Gleason et al., 2018; Holtzer et al., 2008; Watermeyer et al., 2020), thereby relieving clinical and research groups from the pressures surrounding the implementation of novel assessment protocols, such as expertise acquisition through staff training. This may be particularly pertinent to areas of the country or world where resources for such activities are limited (e.g. developing nations). A deeper understanding of the extent to which individual factors, such as personality traits, predict cognitive dispersion may further assist clinicians. That is, while neuropsychologists may consider dispersion to some extent when making a clinical diagnosis, the understanding that inconsistent cognitive profiles are characteristic of individuals high or low in a particular personality trait may better equip medical practitioners to evaluate normative patterns of cognitive variability across individual tests within a neuropsychological battery. For instance, knowledge that individuals higher in neuroticism

are more likely to be characterized by dispersion across cognitive tests may help clinicians to better understand what constitutes normal versus abnormal cognitive dispersion in their patients, potentially providing incremental, but important, information for diagnostic screening or present impairment.

A growing body of research has examined the role of cognitive dispersion as a risk factor for dementia biomarkers and other adverse outcomes, but few studies have focused on factors that may explain individual differences in cognitive dispersion. Personality traits offer a practical option for assessing individuals' tendencies to think, feel, react and behave in a relatively consistent manner across the lifespan (Costa & McCrae, 1992; McCrae & Costa, 2004). While systematic reviews, meta-analyses and coordinated analyses based on extensive reports and samples indicate that personality traits have important implications for cognitive decline (Graham et al., 2021; Luchetti et al., 2016), cognitive complaints (Aschwanden et al., 2020) and risk of dementia (Aschwanden et al., 2021), personality may also have implications for cognitive dispersion in older adulthood. For instance, conscientiousness, characterized by competence, dutifulness and self-discipline (Costa et al., 1991), is positively associated with subjective self-regulation (Reed et al., 2020), while neuroticism, characterized by anxiety, depressive symptoms and emotional instability (Eysenck & Eysenck, 1985), is associated with error-prone performance and impulsivity on measures of executive functioning (Crow, 2019). As the ability to self-regulate and control impulses are likely important for consistent performance across tests within a neuropsychological battery, cognitive dispersion may be characteristic of individuals low in conscientiousness and high in neuroticism.

Additionally, personality traits are related to the experience and perception of stress. High neuroticism contributes to cumulative susceptibility of psychological distress, as well as the associated damaging effects of consistent hypothalamic-pituitary-adrenal axis activation

(Sapolsky, 1996). Likewise, extraversion, which is characterized by sociability, liveliness and activity (Eysenck & Eysenck, 1985), is positively associated with the subjective experience of stress in some studies (Swickert et al., 2002). Individuals high in neuroticism and extraversion may thus demonstrate more variability in performance on a neuropsychological test battery due to test anxiety and emotional instability. Further, as outlined by Curtis et al. (2014), people high in extraversion may perform better on cognitive tasks due to assertiveness, faster responding, and lower general arousal (Chamorro-Premuzic & Furnham, 2004), though individuals high in extraversion may also be more easily distracted and have a lower patience for repetition (Gold & Arbuckle, 1990). Similarly, low conscientiousness and high neuroticism are associated with unhealthy diurnal cortisol patterns, reflecting poor biological coping mechanisms in the face of stress (Montoliu et al., 2020), which may be exacerbated by cognitive testing in older adulthood. Finally, high openness to experience is characterized by desire for and depth of varied emotional experience, as well as cognitive flexibility and intellectual engagement (Costa, 2008; McCrae & Sutin, 2007). Research indicates that mild cognitive impairment (Traykov et al., 2007) and the prodromal stages of vascular cognitive impairment (Garrett et al., 2004) are characterized by poor cognitive flexibility, and that individuals with poor cognitive flexibility are more likely to convert from MCI to dementia (Tatsuoka et al., 2013). As such, openness to experience may contribute to homogeneity in performance across cognitive tests (i.e., less dispersed cognitive performance), as individuals high in openness may approach a neuropsychological battery with cognitive flexibility and receptiveness to cognitive engagement.

Although personality traits are associated with individual differences in cognitive functioning (e.g., (Crowe et al., 2006; Graham et al., 2021; Luchetti et al., 2016)), to our knowledge, no research has examined the extent to which personality traits predict cognitive dispersion, and only one study has investigated the role of personality in cognitive

inconsistency (Munoz et al., 2020). Cognitive dispersion is related to, but distinct from, the concept of cognitive inconsistency, which refers to within-person inconsistencies or fluctuations in performance at repeated attempts at the same task within the same testing occasion (Hultsch et al., 2002). Specifically, Munoz and colleagues (2020) evaluated the role of neuroticism and negative affect in explaining within-person variability across reaction time trials administered 60 times to each participant at the same testing occasion. Munoz et al. (2020) posited that individuals high in neuroticism would have erratic responses in reaction time tasks due to poorer flexibility in emotional and cognitive processes. Consistent with their hypothesis, findings revealed that higher neuroticism was associated with greater variability on repeated reaction time tasks beyond mean reaction time, indicating that neuroticism may be important in the identification and intervention of cognitive dysfunction in older adults.

Munoz et al. (2020) focused on only one personality trait (neuroticism) and variability on repeated administration of only one cognitive functioning test, yet, as postulated above, other personality traits may contribute to variability in performance on cognitive tasks. Previous work indicates that cognitive dispersion is positively associated with cognitive inconsistency, which is expected if variability across various cognitive tests and repeated reaction time tests reflects relatively stable mechanisms (e.g., neurodegeneration) as opposed to dynamic or fluid influences (e.g., pain, fatigue) (Hilborn et al., 2009; Hultsch et al., 2002). Given that clinicians aim to make inferences regarding neurological integrity based on performance across multiple cognitive tests, existing literature, and particularly Munoz et al.'s (2020) findings, justify further exploration of the association between individual differences in personality traits and cognitive dispersion, an index that integrates various cognitive tests and more closely reflects neuropsychological practice.

In the current study, we aim to extend this initial investigation of neuroticism as a predictor of cognitive variability by evaluating the association of all Big Five personality traits and cognitive dispersion across several cognitive measures, drawing data from seven studies of older adults. In an effort to contribute to cumulative science, generate evidence for replicability and generalizability of our research question, and protect against Type I and II errors, we employ a coordinated data analytical approach (Hofer & Piccinin, 2009).

Coordinated analysis is a form of integrative data analysis in which variables are coded consistently across multiple data sets, which are then analyzed independently using the same analytical technique. This approach facilitates the comparison of differences in results based on study-level characteristics (e.g., baseline age of sample, number of cognitive tests), as well as identification of potential patterns of associations across studies. Further, coordinated analysis generates evidence for the replicability and generalizability of a given set of questions. Our pre-registered hypotheses are based on existing literature examining the associations between personality traits and cognitive functioning (Aschwanden et al., 2020, 2021; Baker & Bichsel, 2006; Boyle et al., 2010; Curtis et al., 2014; Duberstein et al., 2011; Soubelet & Salthouse, 2011; Stephan et al., 2021). Specifically, we expect that neuroticism and extraversion will be associated with greater cognitive dispersion; that openness and conscientiousness will be associated with less dispersion; and that agreeableness will not be associated with dispersion.

Methods

Data

Cross-sectional data were drawn from seven international studies of older adults, described briefly below. The measurement occasion in which the Big Five personality traits were first assessed was used in the current analyses. For this project, eligibility criteria required that participants did not have a dementia diagnosis, had data for at least one

personality trait, and had cognitive performance data for at least three cognitive tests. All participants provided informed consent, and ethical approval for each study was granted by governing research committees. Data are available to other researchers by data request via Maelstrom (<https://www.maelstrom-research.org/>). Detailed analytic methods and hypotheses are available via the project preregistration on the Open Science Framework (<https://osf.io/wrnjq/>).

CogUSA

The Cognition and Aging in the USA (CogUSA) study is a longitudinal study of 1,514 adults over the age of 51 living in the United States (McArdle et al., 2015). Data collection took place in three waves between 2007 and 2009. Personality traits were first assessed with the Big Five Inventory (John & Srivastava, 1999) in the second wave, which was selected for analysis. Participants were administered auditory working memory, word recall, number series, picture vocabulary, block design, and the stop/go switch tasks to assess cognition.

ELSA

The English Longitudinal Study of Ageing (ELSA) is a longitudinal study of more than 12,000 English adults over the age of 50 who responded to the Health Survey for England (Stephens et al., 2012). Data collection began in 2002 with additional measurement waves every two years. Personality traits were first assessed in Wave 5, which was selected for analysis, using the Midlife Developmental Inventory (MIDI) (Lachman & Weaver, 1997). Participants were administered word recall, letter cancellation, and verbal fluency tasks to assess cognition.

HRS

The Health and Retirement Study (HRS) is a nationally representative longitudinal panel study of more than 20,000 adults in the United States who were surveyed every two

years starting in 1992 (Sonnega et al., 2014). Personality traits were first assessed in 2006 and 2008, which were selected for analysis, using the MIDI (Lachman & Weaver, 1997).

Participants were administered word recall, numeracy, and backward counting tasks to assess cognition.

LBLS

The Long Beach Longitudinal Study (LBLS) is a longitudinal study of 2,125 adults aged 28-84 living in California (Zelinski et al., 2010). Data collection began in 1978, with an additional six waves added between 1981 and 2008, and two additional cohorts added in 1994 and 2000. Personality traits were first assessed in 1994, which was selected for analysis, using the NEO-PI-R (Costa, 2008). Participants were administered computation span, word recognition, letter series, and verbal fluency tasks to assess cognition.

MIDUS

The Midlife in the United States Study (MIDUS) is a nationally representative longitudinal study of 7,108 adults aged 28-74 (Brim et al., 2004). Data collection began in 1994, with additional waves collected in 2004 and 2013. The Big Five personality traits were assessed at all waves, using the MIDI (Lachman & Weaver, 1997); therefore, we selected the variables collected in 1994 for analysis. Participants were administered digit span, word recall, number series, verbal fluency, and stop/go switch tasks to assess cognition.

SATSA

The Swedish Adoption/Twin Study of Aging (SATSA) is a longitudinal study of 2,019 adults aged 26-93 years. Data collection began in 1984, with additional measurement waves occurring every 3 years (Pedersen et al., 1991). Personality traits were first assessed in 1984, which was selected for analysis, using the NEO-PI (Costa & McCrae, 1985a, 1985b) and EPQ (Eysenck, 1977). Participants were administered digit span, Thurstone picture test, WAIS information test, and digit symbol tasks to assess cognition.

WLS

The Wisconsin Longitudinal Study (WLS) is a longitudinal study of 22,334 Wisconsin residents who graduated from high school in 1957, and their siblings (Herd et al., 2014; Sewell et al., 2003). Personality traits were first assessed in 1992-1994, which was selected for analysis, using the shortened Big Five Inventory (BFI) (John, Donahue, & Kentle, 1991). Participants were administered digit ordering, delayed word recall, number series, and category fluency tasks to assess cognition.

Statistical Approach

Cognitive dispersion scores were derived in each of the studies following an updated formulation of a previously published index of dispersion (Hultsch et al., 2002). The method applies a z-transformation to the raw scores of each test using parameters from the distribution of the entire sample, and then, the application of the formula:

$$Dispersion = \sqrt{\frac{\sum_{k=1}^{K=N} (T_{ik} - S_i)^2}{K-1}}$$

where T_{ik} is the k^{th} test (transformed) for participant i , K is the number of tests, and S_i is participant i 's mean of the transformed scores. Then, linear regression analyses were used to test the association of dispersion scores with each of the personality traits in univariate models, and in conditional models adjusted for mean cognitive performance, education, sex, and age. All variables were z-scored to facilitate interpretation of the results. Education was measured in years in CogUSA, HRS, LBLs, SATSA, and WLS. Education was assessed using ordinal scales in MIDUS (via a 12-point scale) and ELSA (via a 7-point scale), in which higher values represent higher educational attainment. The mean for MIDUS was 7.3 ($SD=2.5$), which indicates three or more years of college, but no completed degree. The mean for ELSA was 4.01 (2.27), which indicates National Vocational Qualification Level 2 (grades A-C on First Diploma in the U.K.). For more information regarding the education variables in MIDUS and ELSA, see

(<https://www.icpsr.umich.edu/web/NACDA/studies/04652/datasets/0001/variables/B1PB1?archive=nacda> and http://doc.ukdataservice.ac.uk/doc/5050/mrdoc/pdf/5050_harmonized_elsa_e.pdf, respectively). As the focus of these analyses was not education, and to promote coordination and facilitate interpretation across studies, all education variables were standardized. Sex was coded as a binary variable (males=0; females=1).

Meta-analyses

Estimates of the associations between dispersion scores and personality traits from each of the studies were meta-analyzed with random effects using the metafor package (Viechtbauer, 2010) in R. The resulting partial correlation coefficients provide an indication of the overall, synthesized association in terms of direction and significance. While this process may obscure meaningful differences between studies to some extent, meta-analysis minimizes Type 1 and Type 2 errors. Further, the random effects approach explicitly addresses between-study variability due to non-identical study characteristics (i.e., no assumption of one true effect size underlying the included studies; instead there may be different true effects for each population) (Hedges & Vellea, 1998). We used the I^2 index (Higgins et al., 2003) to evaluate relative heterogeneity across samples (i.e., the proportion of true variability of the effect relative to the total variability in observed effects) and τ^2 as a measure of between study variance. In particular, I^2 is recommended as a criterion for a decision whether subgroup analysis or moderator analysis is indicated (Borenstein et al., 2009), while τ^2 is used to assign weight to the studies within a meta-analysis under the random effects model. We pre-registered meta-analytic between-study moderator analyses examining the extent to which age and number of cognitive tests included in the computation of cognitive dispersion scores accounted for the association between personality traits and cognitive tests in meta-analytic models indicating substantial heterogeneity.

Openness and Transparency

Following open science recommended practices, we pre-registered our analytical approach and specific hypotheses (URL: osf.io/wrnjq/). Within the pre-registration document, we also report eligibility criteria for participants, inference criteria, all study variables, and links for each of the study datasets. All analyses were conducted in R, and all analysis code and information regarding research materials are available on the open science webpage.

Results

Across studies, baseline characteristics, personality traits, and dispersion scores are reported in Table 1, and ethnic characteristics are reported in Supplementary Table 1. Table 2 reports the mean cognitive test scores and computed cognitive dispersion scores, while Supplementary Table 2 lists the cognitive tests administered in each study and used to compute cognitive dispersion indices across studies. The proportion of men and women differed across studies ($\chi^2 = -73.79, p < 0.001$). Similarly, there were differences in age (ANOVA, $F_{(6,16894)} = 158.5, p < 0.01$) and education (ANOVA, $F_{(5,19578)} = 9167.9, p < 0.01$) across studies.

Dispersion Scores and Personality Traits

Standardized coefficients for personality traits and mean cognitive scores from the fully adjusted linear regression models are reported in Table 3, while the full model results, including p-values and estimates for all covariates, are reported in Supplementary Table 3. Meta-analytic results for the association between dispersion scores and each personality trait are reported in the following subsections.

Neuroticism

Neuroticism was not associated with dispersion in cognitive performance in univariate or in fully adjusted models. Although none of the estimates of the association of neuroticism with cognitive dispersion reached pre-registered significance thresholds (5%), estimates of

the association between neuroticism and cognitive dispersion were negative in four studies (CogUSA, ELSA, LBLs, MIDUS) and positive in three studies (HRS, SATSA, WLS). The pooled effect size from the random effects meta-analysis based on independent analysis of each study was estimated as 0.0017 (SE=0.0056, 95% CI: -0.0092, 0.0126). Heterogeneity estimates indicated no between study variance ($\tau^2 = 0$, SE=0.0001) and I^2 was calculated as 0%, which indicates relative homogeneity between studies. See Supplementary Figure 1 for a graphical representation of the meta-analytic results.

Extraversion

In CogUSA, MIDUS and WLS, positive and statistically significant associations between extraversion and cognitive dispersion in fully adjusted models emerged, suggesting that extroverts have more dispersion in cognitive performance across cognitive tasks. In SATSA the estimate of the association between extraversion and cognitive dispersion scores was also positive, though non-significant, whereas in ELSA, HRS and LBLs, estimates were negative and non-significant (Supplementary Figure 2). The estimated pooled effect of the association of extraversion and cognitive dispersion was 0.0194 (SE=0.0147, 95% CI: -0.0131, 0.0246). τ^2 and I^2 were 0.0003 (SE=0.0003) and 54.67%, respectively.

Openness

The association between openness and cognitive dispersion reached conventional significance levels in HRS, MIDUS and WLS ($\beta=0.014$, SE=0.004; $\beta=0.022$, SE=0.006; and $\beta=0.014$, SE=0.006, respectively) in fully adjusted models. These results indicate that individuals with higher levels of openness had more dispersion in cognitive performance. As these are standardized scores, we can interpret the results accordingly. For example, in MIDUS, individuals who are one standard deviation higher in openness are, on average, 0.02 standard deviations higher in cognitive dispersion. The remainder of the estimates were not statistically significant: two were negative (ELSA, SATSA) while two were positive (LBLs,

CogUSA) (Figure 1). The estimated pooled effect was 0.0285 (SE=0.0129, 95% CI: 0.0030, 0.0537), which was significant at a 10% level. τ^2 and I^2 were 0.0006 (SE=0.006) and 74.58% respectively.

Conscientiousness

The association between conscientiousness and cognitive dispersion reached traditional significance levels only in ELSA, in which conscientiousness was negatively associated with dispersion scores ($\beta=-0.009$, SE=0.004, $p = 0.04$). For CogUSA, LBLs and SATSA, estimates were also negative, whereas in HRS, MIDUS and WLS, estimates were positive, although none of these estimates reached statistical significance levels (Supplementary Figure 3). The estimated pooled effect of the association of conscientiousness and cognitive dispersion was 0.0058 (SE=0.0096, 95% CI: -0.0131, 0.0246). τ^2 and I^2 were 0.0003 (SE=0.0003) and 54.67% respectively.

Agreeableness

In MIDUS and HRS, results revealed significant associations between agreeableness and cognitive dispersion; however, the estimate was positive in MIDUS ($\beta =0.013$, SE=0.006, $p =0.02$) and negative in HRS ($\beta =-0.006$, SE=0.004), $p = 0.01$). Although non-significant, in ELSA, LBLs and SATSA, estimates were also negative, whereas in CogUSA and WLS, estimates were positive (Supplementary Figure 4). The pooled estimate of the association between agreeableness and cognitive dispersion was -0.004 (SE=0.0108, 95% CI: -0.0253, 0.0172). τ^2 and I^2 were 0.0004 (SE=0.0004) and 63.83 % respectively.

Sociodemographic characteristics and cognitive dispersion

Mean cognitive performance was negatively associated with cognitive dispersion in all studies except two, which indicates that individuals with higher mean performance had less dispersed cognitive scores. Mean cognitive estimates were negative and non-significant in LBLs, and positive and non-significant in MIDUS.

A consistent pattern of results did not emerge across studies for the association between any of the sociodemographic characteristics and cognitive dispersion. Age at testing was negatively associated with cognitive dispersion in HRS and SATSA, suggesting that older adults had higher dispersion scores. Baseline age estimates were positive in MIDUS, and significant at $p \leq 0.05$ for all personality traits except openness, indicating that younger adults had higher dispersion scores. Estimates for age at testing were non-significant across the other datasets, but positive in ELSA, negative in WLS and LBLs, and neutral in CogUSA. In MIDUS, the association between sex and cognitive dispersion was positive and significant, suggesting that female participants had higher dispersion scores than male participants. Across the remainder of studies, the association was also consistently positive, but non-significant. In HRS and LBLs, a significant association emerged between education and cognitive dispersion, suggesting that more educated individuals had higher dispersion scores compared to less educated individuals. In the other studies, the associations between education and dispersion were not significant and the estimated direction of the effects were inconsistent (see Supplementary Table 3). Overall, results suggest heterogeneity in the association between sociodemographic variables and cognitive dispersion.

Moderator meta-analyses

We pre-registered study-level moderator tests for the models with substantial heterogeneity to examine if average baseline age (+/- 65 years) or number of cognitive tests included in the computation of cognitive dispersion scores accounted for the association between personality traits and cognitive tests. After preparing the data, however, we realized that comparing the studies based on over or under 65 years old at baseline was not a meaningful comparison, as mean age substantially overlapped across studies, particularly given the deviation in mean age. That is, five of the seven studies were relatively homogenous in terms of age, except for MIDUS, which included relatively young older

adults ($M_{\text{age}}=56.4$, $SD=12.3$), and WLS, which included relatively old older adults ($M_{\text{age}}=71.2$, $SD=0.9$). Therefore, we deviated from the original pre-registration and did not execute moderator analyses examining the impact of mean age. For heterogeneous models (all models except neuroticism), we executed moderator meta-analyses examining the impact of number of cognitive tests included in the computation of cognitive dispersion. Results revealed that the number of cognitive tests used in the computation of cognitive dispersion did not moderate the association between personality and cognitive dispersion.

Discussion

The current study examined five pre-registered hypotheses focused on the relationships between personality traits and dispersion in cognitive performance applying a coordinated analysis approach to data from seven studies of older adults. We postulated that neuroticism and extraversion would be associated with more dispersion, whereas openness and conscientiousness would be associated with less dispersion, and no association would exist between agreeableness and dispersion. Results from the random effects meta-analyses showed that the only statistically significant pooled estimate was a positive association between openness and cognitive dispersion, which was inconsistent with our expectations based on existing literature. Overall, these findings suggest weak evidence in support of our predictions. Analyses based on individual studies, however, revealed some significant associations between individual differences in personality traits and cognitive dispersion. Specifically, consistent with our predictions, results revealed significant positive estimates between extraversion and cognitive dispersion in three out of seven associations, but the overall, meta-analytic estimate was non-significant. Further, though conscientiousness and cognitive dispersion were negatively associated in one study, the remainder of results revealed non-significant associations in inconsistent directions. Finally, consistent with our expectation that no association would emerge between agreeableness and cognitive

dispersion, only one study revealed a significant association, though these findings were similar to the inconsistency observed across the other traits. While findings based on individual study results may provide direction for future research, we focus our discussion on synthesized results based on the meta-analyses.

As mentioned, synthesized results revealed a positive association between openness and cognitive dispersion for individuals of the same age and gender, with the same education and average cognitive performance, suggesting that individuals with higher openness scores have higher cognitive dispersion. The meta-analytic estimate in the opposite direction warrants careful interpretation, as we pre-registered the hypothesis of a negative association between openness and cognitive dispersion. We predicted that the characteristics of individuals high in openness may contribute to a better ability to shift between cognitive tests (e.g., creativity, cognitive flexibility and receptiveness to cognitive engagement). Moreover, existing literature indicates that higher openness is associated with better cognitive functioning (Curtis et al., 2014; Luchetti et al., 2016; Soubelet & Salthouse, 2011). Nevertheless, we cautiously consider these findings. It is possible that, rather than high openness leading to more flexibility between cognitive tasks, individuals higher in openness are more engaged by cognitive tasks that require more flexibility and creativity. For instance, given the tendency to be more imaginative and gravitate towards varied emotional experience, individuals high in openness may thrive on cognitive tasks that elicit more creativity (e.g., verbal fluency, word recognition) or perspective shifting (e.g., digit symbol, figure rotation), relative to cognitive assessment that requires memorization or working memory (e.g., number series, computation span, backward counting).

We did not, however, consider the shape of cognitive dispersion profile within these analyses; as such, two individuals could have quite heterogeneous cognitive profiles despite identical cognitive dispersion scores. Future research examining the association of facets of

openness with cognitive dispersion, as well as shape of cognitive dispersion profiles, may provide a more nuanced understanding of the mechanisms underlying these findings and opportunities for application within clinical settings. For instance, in the context of a substantial association between personality traits (or trait facets) and cognitive dispersion, clinicians may consider including a personality test to bolster their assessment of health and risk stratification. Within the current analyses, the strength of the association between openness and cognitive dispersion was small, though a variety of factors likely influence cognitive inconsistency, impairment, or decline in older adulthood; thus, any individual predictor is likely to demonstrate only a small-to-medium relationship. Future work could explore this association relative to more established risk factors for cognitive impairment, such as genetic, health, and lifestyle indicators, in order to guide clinical interpretations of patients' cognitive dispersion metrics alongside personality parameters.

The meta-analytic models revealed heterogeneity levels that ranged from small (e.g., neuroticism, $I^2=0\%$) to substantial (openness, $I^2=74.58\%$). We executed meta-analytic between-study moderator analyses to examine if the number of cognitive tests used in the computation of cognitive dispersion moderated the association between personality traits and cognitive dispersion in models with substantial heterogeneity (all models except neuroticism). Results revealed that the number of cognitive tests used in the computation of cognitive dispersion did not moderate the association between personality and cognitive dispersion across any trait. We encourage follow-up investigation in this area, however, as I^2 may be biased in meta-analyses based on a small number of independent effects (von Hippel, 2015).

Personality traits as predictors of individual differences in cognitive dispersion is an emerging area of inquiry, which limits our ability to compare our findings with existing reports. Munoz et al. (2018) found that neuroticism predicted greater reaction time variability

(i.e., cognitive inconsistency) across ages, independent of mean response time and demographic covariates. Our results failed to expand upon these findings in our examination of the association between cognitive dispersion and neuroticism in the overall, synthesized results, or in the independent analyses across seven samples. Interestingly, the estimates between neuroticism and cognitive dispersion were negative but not significant in four of the seven studies, suggesting that individuals with lower levels of neuroticism may have more dispersion in cognition than individuals with higher levels of neuroticism. These results are in partial contradiction with the results reported by Munoz et al. (2018), though our findings are not directly comparable as cognitive dispersion indices measure a different construct than reaction time variability. Yet, given previous evidence showing that individuals with higher levels of neuroticism are more error prone while completing cognitive tasks (Robinson et al., 2006), we encourage further investigation of potential explanations for these findings.

We found some evidence of a negative association between mean cognitive performance and cognitive dispersion, which indicates that individuals with higher average cognitive performance had less dispersion in performance across cognitive tasks. Across the majority of studies and models, sociodemographic variables were not significantly associated with cognitive dispersion; further, mixed results emerged regarding the direction of trends. Previous investigations of intra-individual variability in reaction time have generally shown increased variability in older participants and in participants with lower mean cognitive performance (Bielak et al., 2010). However, as previously mentioned, cognitive dispersion captures a distinct construct compared to cognitive inconsistency in reaction time.

It is worth noting that compared to the other studies, MIDUS had the most significant associations between personality traits and cognitive dispersion (openness, extraversion and agreeableness were positively associated with cognitive dispersion). Although the average dispersion scores in MIDUS did not differ significantly from average dispersion in other

studies, MIDUS was also the only study where mean cognitive performance was positively associated with cognitive dispersion, albeit a statistically non-significant association.

Interestingly, the cognitive battery and individual cognitive tests available in MIDUS overlapped substantially with other studies; for instance, MIDUS administered identical cognitive tests (digit span, word recall, number series, Stop/Go switch task, and verbal fluency) compared to the other included studies. As average dispersion in MIDUS did not differ from other studies, and due to similarity in administration of cognitive tests compared to the other studies, differences in the MIDUS results may be due to other between-study differences in study characteristics. For instance, MIDUS is the youngest of the cohorts considered here ($M_{\text{age}}=56.4$, $SD=12.3$), and participants have the fewest years of education ($M_{\text{educ}}=7.28$, $SD=2.54$).

Importantly, age at testing was significantly and positively associated with dispersion scores in MIDUS, whereas age at testing was significantly and negatively associated with cognitive dispersion in HRS and SATSA, which both include relatively older participants. Given that participants in MIDUS were on average a decade younger than participants in HRS and SATSA, these results indicate that in a sample of relatively younger adults, the youngest adults tend to have higher dispersion scores, whereas in a sample of relatively older adults, the oldest adults have higher dispersion scores. Together, these results point to the possibility of a non-linear, U-shaped relationship between cognitive dispersion and age; specifically, cognitive dispersion may be more substantial in young-old adults (better performance on some tests) and old-old adults (worse performance on some tests), while middle-old adults may tend to regress to their mean. These findings are consistent with research showing that age significantly impacts the association between cognitive dispersion and neurofibrillary tangles (NFT) (Malek-Ahmadi et al., 2017), such that cognitive dispersion is more strongly related to NFT for those dying at younger ages. Likewise, existing research

suggests that educational attainment may influence cognitive dispersion in middle-old adults (~65 years old), but not in late-old adults (~80- 90 years old), suggesting more efficient compensatory strategies in response to cognitive or neuronal senescence in younger age groups (De Felice & Holland, 2018; Watermeyer et al., 2020, 2021). We encourage researchers to extend our investigation to examine age-related compensatory strategies that may occur in response to cognitive senescence.

This investigation includes both strengths and limitations. The independent analyses conducted using seven large established studies of older adults substantiate the limited findings, while the synthesis of results further enhance the current research. In addition, we preregistered the project, including the hypotheses, on the Open Science Framework, which contributes to transparency of research. The only deviation from our original plan was not executing a meta-analysis examining age (over/under $M_{age} = 65$ years) as a between-study moderator, since mean age was clustered around 65 years in five of the studies included. Furthermore, utilizing a coordinated analysis approach, we maximized the use of all available data from cognitive tests in each of the studies, rather than coordinating at the lowest possible denominator. As such, there were between-study differences in the measures included in the derivation of cognitive dispersion. However, results remained relatively consistent across studies (i.e., limited evidence for an association between cognitive dispersion and individual differences in personality traits), despite differences and similarities in the individual tests included in the derivation of dispersion scores. Previous studies also vary in the measures included in the derivation of the scores, which suggests the index may be robust to these differences. Nevertheless, future research could consider the association between personality traits and the shape of cognitive dispersion profile, similarly to Peters et al. (2005).

The included studies also varied in several key characteristics, including the personality tests that were administered and the age at which personality was first assessed,

which subsequently impacted the age at which we investigated the association between personality traits and cognitive dispersion. Another possible limitation is that we restricted our analyses to personality and cognitive dispersion, potentially neglecting other important variables, such as health (or disease) factors, which might contribute to inconsistency across tasks and level of cognitive performance, as well as correlate with personality domains (Strickhouser, Zeel & Krizan, 2017). Unfortunately, this was due in large part to inconsistencies in health assessments across the studies, which were selected on the basis of prioritizing the harmonization of personality measures and cognitive domains. Finally, the included studies are based in industrialized countries, the majority of participants are highly educated, and 83%-100% of participants identified as White or Caucasian across the seven studies (see Supplementary Table 1). As such, our results may be limited to generalizing to predominately highly educated and industrialized Caucasian populations. Research based on personality surveys in 23 low- and middle-income countries ($N=94,751$) suggests that assessment of the Big Five personality traits may not validly capture the intended personality traits outside of Western, education, industrialized, rich, and democratic populations (Laajaj et al., 2019). Follow-up research focused on diverse samples that accounts for potential cross-cultural differences in interpretation of trait scales would benefit the existing literature.

To our knowledge, this investigation is the first to examine the relationship between personality traits and cognitive dispersion. It builds upon growing evidence that supports cognitive variability as a marker for cognitive and neurological dysfunction by attempting to delineate the influence of individual differences in personality traits on dispersion scores across several older-age groups. Apart from age and education, there has been limited exploration of other variables that may influence cognitive variability. We encourage researchers to further examine the associations between personality traits and cognitive dispersion longitudinally and at different times during older adulthood, as it is possible that

associations emerge as individuals' cognitive functioning deteriorates and inconsistency in performance across tests becomes more substantial.

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Table 1. Baseline Characteristics and Personality Traits across Studies

Study	N	Female	Male	Age	Education	Neuroticism	Openness	Conscientiousness	Extraversion	Agreeableness
CogUSA	1207	647 (56%)	533 (44%)	64.09 (10.65)	14.19 (2.35)	37.27 (20.57)	68.05 (17.78)	77.44 (15.25)	59.08 (20.45)	81.77 (13.41)
ELSA	8771	4907 (56%)	3864 (44%)	66.68 (10.04)	4.01 (2.27)	2.59 (0.47)	2.12 (0.56)	1.60 (0.51)	1.85 (0.56)	1.49 (0.48)
HRS	14863	8862 (60%)	6001 (40%)	67.22 (11.30)	12.84 (3.03)	2.03 (0.62)	2.91 (0.58)	3.36 (0.50)	3.17 (0.58)	3.51 (0.50)
LBLS	590	307 (52%)	283 (48%)	69.45 (13.92)	13.81 (3.01)	79.43 (20.78)	105.79 (17.65)	121.27 (17.64)	104.27 (18.32)	125.09 (15.92)
MIDUS	3671	2029 (55%)	1642 (45%)	55.43 (12.45)	7.3 (2.5)	2.07 (0.63)	2.90 (0.54)	3.46 (0.45)	3.11 (0.57)	3.45 (0.50)
SATSA	625	385 (62%)	240 (38%)	65.53 (8.39)	NA	2.82 (2.32)	2.99 (0.43)	3.84 (0.48)	4.75 (2.26)	3.92 (0.44)
WLS	3855	2087 (54%)	1768 (46%)	71.27 (0.95)	13.6 (2.3)	14.82 (4.71)	20.25 (4.90)	27.96 (4.75)	22.22 (5.49)	28.28 (4.82)

Note. Values are (sub)sample size and %, or means and standard deviations; CogUSA= The Cognition and Aging in the USA; ELSA=English Longitudinal Study of Aging; HRS=Health and Retirement Study; LBL=Long Beach Longitudinal Study; MIDUS=Midlife in the United States; SATSA=Swedish Adoption Twin Study of Aging; WLS=Wisconsin Longitudinal Study; **Age and education in years, except for education in ELSA and MIDUS, which were measured according to ordinal scales variables ranging 1-7 and 1-12, respectively, in which higher values indicate high educational attainment.**

Table 2. Average cognitive test scores and computed dispersion scores across studies

Study	Cognitive Tests, Aligned According to Cognitive Domains							Computed Cognitive Dispersion Score
	Working Memory	Declarative Memory	Fluid Intelligence	Crystallized Intelligence	Visuospatial Reasoning	Speed	Verbal Fluency	
	<i>M (SD)</i>	<i>M (SD)</i>	<i>M (SD)</i>	<i>M (SD)</i>	<i>M (SD)</i>	<i>M (SD)</i>	<i>M (SD)</i>	
CogUSA	516.73 (22.36)	11.11 (3.20)	55.53 (10.79)	55.55 (11.10)	51.25 (9.92)	534.65 (28.25)		0.70 (0.27)
ELSA		4.59 (2.16)				4.26 (1.95)	5.75 (2.25)	0.75 (0.40)
HRS		9.75 (3.25)	508.34 (40.34)			1.87 (0.50)		0.72 (0.54)
LBSL		17.77 (2.44)	11.27 (6.75)		18.29 (10.97)		33.23 (12.53)	0.71 (0.35)
MIDUS	4.96 (1.53)	4.36 (2.63)	2.18 (1.53)			-1.10 (0.29)	18.85 (6.17)	0.77 (0.34)
SATSA	9.68 (2.15)	20.60 (4.62)		30.91 (8.41)	18.44 (7.70)	38.38 (12.29)		0.71 (0.27)
WLS	5.63 (1.42)	3.44 (1.79)	8.39 (3.67)				11.15 (4.17)	0.82 (0.36)

Note. All reported values are raw scores, except Speed for MIDUS, which is transformed so that higher values indicate better cognitive performance (these values were standardized prior to analysis). See Table 1 for study title abbreviations; *M*=Mean; *SD*=Standard Deviation; See study description in text or Supplementary Table 1 for specific cognitive tasks administered aligned according to cognitive domains.

Table 3. Personality trait and mean cognitive score estimates from multivariable linear regression models fitted to cognitive dispersion scores in each of the seven international studies adjusted for age, sex, education, mean cognitive scores, and personality traits.

Study	CogUSA	ELSA	HRS	LBLS	MIDUS	SATSA	WLS
	<i>Beta (SE)</i>						
Cog dispersion							
Neuroticism	-0.002 (0.007)	-0.001 (0.004)	0.003 (0.004)	-0.008 (0.015)	-0.001 (0.006)	0.002 (0.011)	0.003 (0.006)
Mean Cog Score	-0.143 (0.014)**	-0.011 (0.007)	-0.495 (0.006)**	-0.052 (0.028)	0.015 (0.010)	-0.076 (0.016)**	-0.045 (0.010)**
Cog dispersion							
Extraversion	0.015 (0.007)*	-0.007 (0.004)	-0.002 (0.003)	-0.0001 (0.015)	0.020 (0.006)**	0.013 (0.011)	0.013 (0.006)*
Mean Cog Score	-0.143 (0.014)**	-0.014 (0.007)*	-0.496 (0.006)**	-0.050 (0.028)	0.015 (0.010)	-0.076 (0.016)**	-0.045 (0.010)**
Cog dispersion							
Openness	0.012 (0.008)	-0.006 (0.004)	0.014 (0.004)**	0.006 (0.016)	0.022 (0.006)**	-0.008 (0.012)	0.014 (0.006)*
Mean Cog Score	-0.146 (0.014)**	-0.015 (0.007)*	-0.497 (0.006)**	-0.052 (0.028)	0.014 (0.010)	0.014 (0.006)**	-0.05 (0.01)**
Cog dispersion							
Conscientiousness	-0.002 (0.007)	-0.009 (0.004)*	0.006 (0.004)	-0.005 (0.015)	0.010 (0.006)	-0.004 (0.011)	0.007 (0.006)

Mean Cog Score	-0.142 (0.014)**	-0.014 (0.007)*	-0.497 (0.006)**	-0.050 (0.028)	0.015 (0.010)	-0.067 (0.017)**	-0.046 (0.010)**
Cog dispersion							
Agreeableness	0.000 (0.007)	-0.008 (0.004)	0.006 (0.004)*	-0.016 (0.015)	0.013 (0.006)*	-0.001 (0.012)	0.003 (0.006)
Mean Cog Score	-0.142 (0.014)**	-0.012 (0.007)	-0.497 (0.006)**	-0.049 (0.028)	0.015 (0.010)	-0.068 (0.017)**	-0.045 (0.010)**

Notes. Estimates are reported as standardized coefficients; See Table 1 for study title abbreviations; SE=Standard error; Mean Cog Score=Mean cognitive score; See Supplementary Table 2 for all estimates from the models (trait, mean cognitive score, age, sex, and education) and all p-values; *p<0.05, **p<0.001

Figure 1. Meta-analytic results for the partial correlation coefficient between openness and cognitive dispersion across studies.