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1 Gender differences and mask wearing: an observational study on a University campus and a
2 mini-meta-analysis

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5 Author Note

6 Author note to be included.

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Abstract

10

11 Research informed by evolutionary theory has suggested that, all else being equal, men are
12 expected to take greater risks than women. This has been evidenced in a range of domains,
13 including health prevention behaviours. In this study, gender differences in mask wearing
14 were recorded at three locations on a University campus ($n = 1,435$). Logistic regression and
15 Bayes Factor analyses demonstrated that the data do not support a gender difference in
16 mask wearing. This led us to supplement our findings with a mini-meta-analysis, synthesising
17 the gender difference reported in ten papers ($n = 73,493$) observing mask wearing during the
18 COVID-19 pandemic. This analysis is supportive of a weak effect whereby women are more
19 inclined to wear a mask than men ($OR = 1.54$, $95\% CI = 1.26$ to 1.88). However, the
20 mini-meta-analysis also suggested a considerable amount of heterogeneity. Our research calls
21 for further work assessing the factors explaining this heterogeneity in the observed gender
22 difference in mask wearing. **This is a preprint, currently under consideration at a**
23 **journal. This version 01-11-2022**

24

Keywords: Risk Taking; Observation; Health behaviour; COVID-19; Gender

25

Word count: 3,733

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27 mini-meta-analysis

28 Evolutionary research has put forward that there are evolved differences between men
29 and women (e.g., Buss, 2019), including in the behavioural tendency to take risks (e.g., Daly
30 & Wilson, 2001; Wilson & Daly, 1985). While the size of this difference can vary based on
31 age, culture and contextual factors, research from a range of domains supports the existence
32 of a difference in risk taking (e.g., Finance: Olsen & Cox, 2001; Economic games: Charness
33 & Gneezy, 2012; Psychological scales: Weber, Blais, & Betz, 2002). Meta-analyses on risk
34 taking attitudes also provide support for the existence of a gender difference, though the size
35 of the effect varies between domains (e.g., Byrnes, Miller, & Schafer, 1999). These gender
36 differences in risk taking also play out in every day decisions. This is clear in traffic
37 behaviour: men are less likely to wear a seat belt than women (e.g., Calisir & Lehto, 2002;
38 Lerner et al., 2001). Men are also less likely to use lights on their bicycle at night than
39 women (Cobey, Laan, Stulp, Buunk, & Pollet, 2013). Men are also more likely to unsafely
40 cross the road than women (e.g., Pawlowski, Atwal, & Dunbar, 2008; Pollet & O'Dowd,
41 2018).

42 Gender differences are also evident in both health risk taking behaviour and
43 preventative behaviours. For example, all else being equal, Pinkhasov et al. (2010) found
44 that American men are more likely to be regular and heavy alcohol drinkers, heavier
45 smokers, and illegal drug users, compared to American women. Men were also found to be
46 less likely to utilize health care than women: for example, visiting doctor's offices or going to
47 emergency departments. Men were also less likely to make use of preventative care or dental
48 care visits than women. In a wide variety of domains, these gender differences in
49 preventative health behaviours have been documented (for example: screening for skin
50 cancer: Evans, Brotherstone, Miles, & Wardle, 2005; diabetes management: Shalev, Chodick,
51 Heymann, & Kokia, 2005), but note that exceptions do exist (for example: methods for

52 colorectal screening: McMahon Jr et al., 1999). All else being equal, it is therefore no
53 surprise that the health prevention literature leads to argue that men and women will behave
54 differently when it comes to prevention during a pandemic (e.g., review in context of H1N1:
55 Bish & Michie, 2010).

56 In the context of the COVID-19 pandemic, gender differences in preventative
57 behaviours have indeed been consistently documented. Even though data suggest that men
58 are at greater risk than women (e.g., Rushovich et al., 2021), men appear to be less inclined
59 to protect themselves from COVID-19 than women. For example, multiple studies using
60 surveys indicated that men are less likely to take preventative measures against COVID-19
61 than women (e.g., Hearne & Niño, 2022; Latkin et al., 2021; Padidar et al., 2021). These
62 measures including willingness to wear a mask and thus whether or not to wear a mask is an
63 every day decision in a health context entailing risk. Systematic reviews provide evidence
64 that, physical barriers, i.e. masks, reduce the risk of transmission of airborne viruses (e.g.,
65 Chu et al., 2020; Jefferson et al., 2009; Liang et al., 2020). When COVID-19 took hold of
66 countries, governments, under the advice public health officials, therefore implemented mask
67 requirements (e.g., in Belgium, France, United Kingdom, Badillo-Goicoechea et al., 2021).
68 When the ‘Omicron’ variant became dominant in the UK, masks became a [requirement](#) for
69 public transport, shops, and were also recommended for universities.

70 Compared to the volume of survey research on preventative behaviours toward
71 COVID-19, there have been relatively few studies of mask wearing relying on direct
72 observation. A study capturing data from live streaming of high school graduations in 5 US
73 high schools (n = 1,152), found that nearly 70% of students wore a mask, but did not find a
74 gender difference. A small study, covering three locations in North-East of the U.S., found
75 that women were more inclined to wear masks than men (n= 300, Okten, Gollwitzer, &
76 Oettingen, 2020). Haischer et al. (2020) conducted observations at retail locations in
77 Milwaukee (US) in June/July 2020 (n = 9,935). These authors found that the odds of mask

78 wearing was greater for women than for men by around a factor of 1.5. An observational
79 study of 1,004 people in Vermont (US) also found that women were more inclined to wear
80 masks than men (Beckage, Buckley, & Beckage, 2021), as did a study in New York city (US)
81 parks (n = 1,453, Hitch et al., 2022). Data from a range of non-Western countries similarly
82 support a gender difference in mask wearing, for example data from Argentina (n = 15,507,
83 Freidin, Acera Martini, Senci, Duarte, & Carballo, 2022), Taiwan (n = 11,680, Chuang &
84 Liu, 2020) and Ethiopia (n = 632, Woldearegay, 2022).

85 **Current study**

86 Given that the majority of research suggests a gender difference in mask wearing, the
87 predictions are that women would be more likely to wear masks (1) and more likely to use
88 hand sanitiser (2) than men. The hand sanitiser was available from a stand at each building
89 but during our study not a single participant was observed using hand sanitiser. Therefore
90 this measure does not feature in this paper. No predictions were made about location on
91 campus or the interaction between gender and location on mask use. These analyses are thus
92 exploratory.

93 **Methods**

94 **Observations**

95 The observation schedule and analysis plan was preregistered. The sample size was
96 determined by the time allocated to the first author for this project (data collected between
97 Feb. 1st and Feb. 25th). There were three locations: Library, Student union, Gym, at a large
98 university campus in England. Two data collection sessions were planned at each location.
99 One session was rescheduled due to inclement weather. The procedure was approved by the
100 local ethics committee where the study was carried out. The University had a non-enforced
101 requirement for staff and students to wear a face mask when moving in buildings on campus.
102 Two raters coded perceived gender (male/female), mask use (yes/no) and hand sanitation

103 use (yes/no) when entering or exiting a campus building. Any type of face covering, and
104 regardless of how it was worn, was coded as mask use. Groups of individuals were coded
105 individually and great care was taken not to code the same individual twice, if they were, for
106 example, to exit after a short stay. The first twenty observations were used to trial the
107 procedure and discarded. There was perfect agreement between the raters on mask wearing
108 but three cases where the codes did not correspond for gender. These were likely coding
109 errors and were excluded. This leads to a final sample of 1,435 observations. It should be
110 noted that just prior to the study, [on January 26th](#), the UK government announced the
111 removal of the requirement for wearing masks indoors. However, the Mayor for London
112 announced that they would still [be required for the London metro](#). This continued
113 requirement to wear a mask also applied to the University where the work was carried out:
114 mask wearing remained compulsory when navigating through any campus building and this
115 was communicated to staff and students.

116 **Data analysis**

117 All the analyses were conducted in R 4.2.1 (R Development Core Team, 2008). The
118 data, code, and analysis document are available from the [Open Science Framework](#). The key
119 hypothesis test is evaluated with a logistic regression model. An *a priori* power analysis via
120 G*Power 3.1, suggested that a sample size of $n = 192$ is required to detect a weak effect
121 (Odds Ratio of 1.68, Chen, Cohen, & Sophie Chen, 2010) at 80% power and a two-tailed
122 significance level of 5% (Faul, Erdfelder, Lang, & Buchner, 2007). Next to frequentist
123 statistics, Bayes Factors (BF) which allow comparing models (Makowski, Ben-Shachar, &
124 Lüdtke, 2019) are presented. Many rules of thumb for the interpretation of BFs exist
125 (Jarosz & Wiley, 2014). Here, the qualifications for evidence by Jeffreys (1961) were used
126 (BF = 1 - No evidence, $1 < BF \leq 3$ - Anecdotal, $3 < BF \leq 10$ - Moderate, $10 < BF \leq$
127 30 - Strong, $30 < BF \leq 100$ - Very strong, $BF > 100$ - Extreme).

128

Results

129 Figure 1 represents the data. Even though masks were compulsory based on university
 130 guidelines, in only 28.6% of the observations the person was wearing a mask.

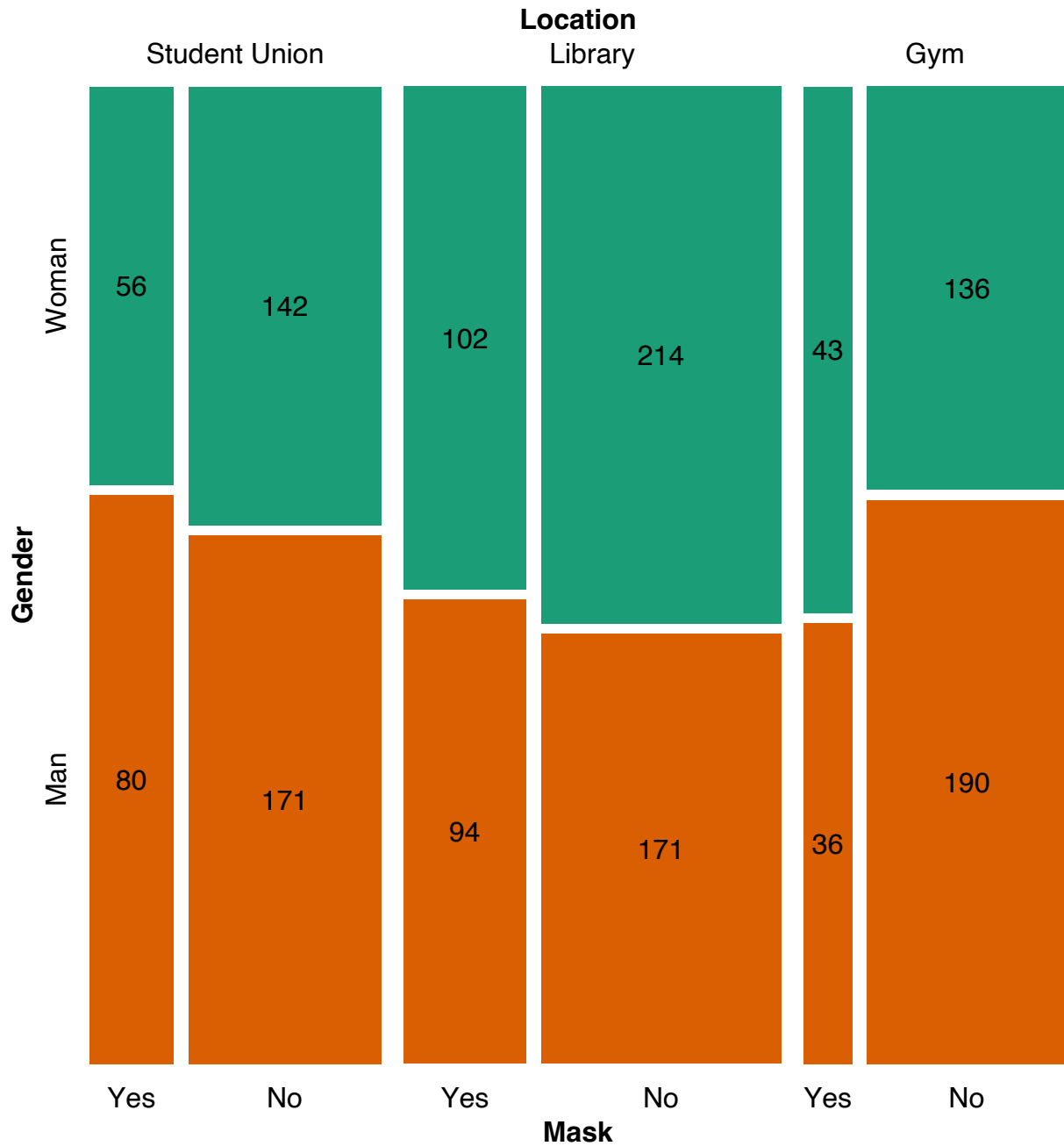


Figure 1. Mosaic plot for gender, location and mask wearing.

131 Table 1 shows the results from logistic regression models. Model 1 contains Gender,
132 Model 2 contains Location, Model 3 contains Gender and Location, Model 4 contains the
133 main effects of Gender and Location and an interaction effect. Model 1 showed that there is
134 no support for gender differences in mask wearing ($\chi^2(1) = .086, p = .769$). Model 2 showed
135 that individuals were more likely to wear a mask when entering or exiting the library and
136 student union compared to the gym. For the Library 33.73% of individuals were recorded to
137 be wearing a mask, for the Student union 30.28% , as opposed to 19.51% for the gym. Model
138 3 showed that the location effect is upheld, when including gender in the model. Model 4
139 suggested an interaction effect between gender and location on mask use. As demonstrated
140 in Figure 1, women were more likely than men to wear a mask at the gym. However, while
141 some of the individual coefficients were statistically significant, the likelihood ratio test for
142 the interaction effect was not ($\chi^2(2) = 5.47, p = .065$). As this result could be considered on
143 the cusp, and given that there was not an *a priori* predicted an interaction effect, Bayes
144 Factors which allow quantifying the evidence for one model versus another were also used.

Table 1

Odds Ratios for logistic regression models predicting mask wearing. Note: Reference categories are: "Man" and "Gym".

| | Outcome | | | |
|---------------------|-----------|-----------|-----------|-----------|
| | (1) | (2) | (3) | (4) |
| Woman | 1.035 | | 0.993 | 1.669* |
| Student union | | 2.101*** | 2.102*** | 2.901*** |
| Library | | 1.793*** | 1.793*** | 2.469*** |
| Woman*Student union | | | | 0.520* |
| Woman*Library | | | | 0.505* |
| <i>N</i> | 1,435 | 1,435 | 1,435 | 1,435 |
| Log Likelihood | -859.388 | -846.641 | -846.639 | -843.904 |
| AIC | 1,722.776 | 1,699.282 | 1,701.279 | 1,699.807 |

*p < .05; **p < .01; ***p < .001

145 Bayes Factors very strongly favour the null model by a factor of >36 over the model
146 with gender (Model 1). The null model is also strongly favoured over the model with the
147 interaction by a factor >14 (Model 4). In contrast, the location model (Model 2) is favoured
148 by a factor >250 over the null model. Finally, if the analysis is restricted to the Gym
149 location, then the null model is still favoured by a factor of 2.54 over a model with gender
150 included, albeit this is only ‘anecdotal’ evidence for the null model. In sum, even in a
151 subgroup analysis, the data do not support a gender effect on mask wearing.

152 **Mini-meta-analysis**

153 Effect sizes on the gender difference in mask wearing uncovered in the above literature
154 review were synthesised in a mini-meta-analysis. Only peer-reviewed papers were included.
155 Papers needed to rely on direct observation of (adult) mask wearing during the COVID-19
156 pandemic and needed to allow for derivation of an odds ratio for gender. The log(odds ratio)
157 was synthesised via a random effects meta-analysis with restricted maximum likelihood
158 (REML, Viechtbauer, 2010). More details and additional analyses can be found on [the OSF](#).

159 Figure 2 demonstrates that, overall, there is support for a weak effect of gender: an
160 odds ratio of 1.54 (95% CI: 1.26, 1.88). It also demonstrates a very large heterogeneity
161 between the studies, with individuals studies ranging from no support to an odds ratio of
162 2.47 (95% CI: 1.82 to 3.35).

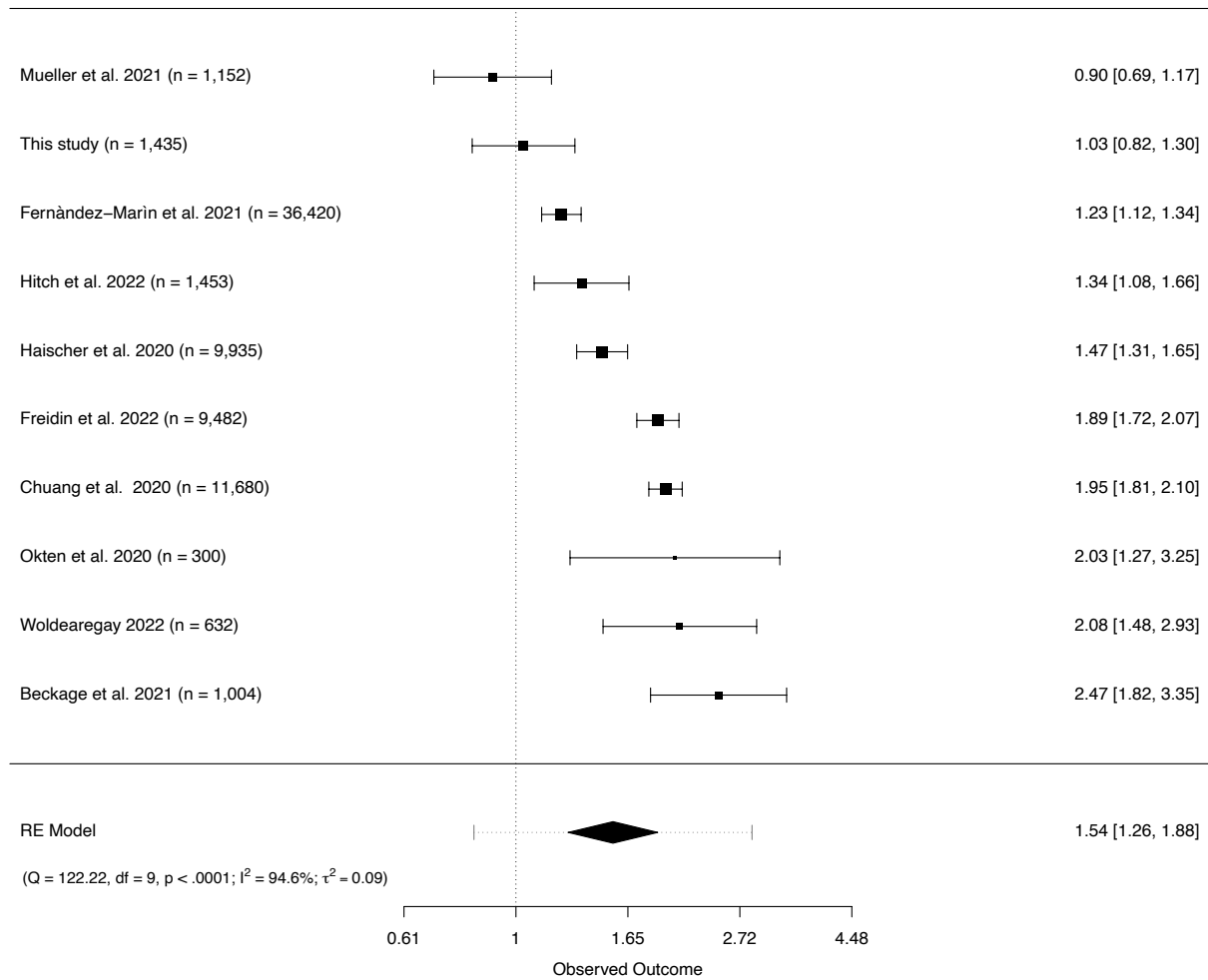


Figure 2. Forest plot of odds ratios (back transformed from meta-analysis for figure). The tips of the diamond present the 95% CI for the pooled effect size estimate, the dashed interval represents the prediction interval.

Discussion and conclusion

163

164 For the observational study, Bayesian analyses actively support the null model: the
165 data did not support the key prediction that women would be more inclined to wear a mask
166 on campus. Also the data showed that mask wearing, even though compulsory on University
167 campus, was not the norm. Less than 1 in 3 individuals were observed to be wearing a face
168 covering. This could be one potential reason for why our study did not find a gender
169 difference: Mask wearing was not the norm. However, in the setting where masks were least
170 normative (Gym), the gap between men and women in mask wearing was the largest.
171 Therefore, this explanation is perhaps unlikely.

172

173 There are many limitations to the current observational study. First, behaviour was
174 recorded at only one university. However, we would expect similar findings at other
175 universities in the UK. It is unclear how our campus would be (very) different from others,
176 though campuses could vary in gender composition and the degree to which mask wearing
177 was the norm. Second, gender was inferred via observation and, even though there was near
178 perfect agreement on perceived gender, this is a clear limitation of any observational study
179 on gender differences. Though there is bound to be some error in this variable, it seems
180 unlikely that this would overturn the strong evidence in favour of the null model, as opposed
181 to the model containing gender (Bayes Factor >36). Third, as individuals were only
182 observed on campus, and given that social isolation is a risk mitigating strategy, it is possible
183 that men were more likely to come on to campus than women. This is not something that
184 can be ruled out based on the collected data, as it would require estimating the number of
185 individuals which would normally be at the campus locations. Moreover, it is also possible
186 that women were taking additional health protective measures which were not recorded, such
187 as using their own hand sanitiser, keeping their distance, or washing their hands more
188 frequently, compared to men. Therefore, it is possible that if a wider range of preventative
behaviours is examined there would be support for a gender difference in preventative

189 behaviours. Fourth, even though our measure of wearing masks is a reasonable proxy for a
190 health preventative behaviour, it is unclear which psychological mechanisms actually
191 influence mask wearing. While mask wearing was not normative in our setting, the motives
192 for wearing or not wearing a mask could be shaped by conformity more so than
193 considerations relating to health risk. Thus, this study does not allow disentangling whether
194 individuals wore masks due to compliance or for health reasons. Conversely, it is unclear if
195 not wearing a mask would constitute non-compliance or risk taking. Further research on
196 behavioural intentions is needed to better understand why individuals choose to wear masks
197 and the potential role which gender might play on the context in which one is worn.

198 Our mini-meta-analysis demonstrated that across ten observational studies there was
199 some support for a weak gender difference in mask wearing, with women being more inclined
200 to wear a mask than men. Nonetheless, these analyses also demonstrated substantial
201 heterogeneity in this effect (Figure 2). This is evident from the wide prediction interval, for
202 example. The prediction interval includes an odds ratio of 1: a new study could thus still be
203 expected to find an odds ratio of 1, i.e. no difference. Our research calls for more
204 observational work in this area, as well as research into the factors which could explain the
205 observed heterogeneity (e.g., how widespread the norm is of mask wearing, prevalence of risk,
206 culture, overall gender composition of the sample, etc.). As described above, future work is
207 also necessary to examine the mechanisms which could lead to the observed gender
208 difference. For now, even though our individual study did not support a gender difference,
209 we conclude that the available data are consistent with a weak, but heterogeneous, observed
210 gender difference in mask wearing.

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